ON THE USE OF LATIN SQUARES FOR DESIGNING CARRY-OVER EXPERIMENTS

A THESIS

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No man ought to be discouraged if the experiments he puts in prectice answer not his expectations, for what succeeds pleases more, but what succeeds not many times, informs no less. "BACON"



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PREFACE

The thesis aims at utilizing latin squares in the design and analysis of carry-over effects of treatments in block experiments. It also aims at preparing computer programs for the processing of data in some of the designs presented in the thesis.

The thesis is composed of three parts, the first of which introduces general mathematical approach for block experiments with estimators of parameters in the linear model adopted for these experiments and the types of statistical tests of linear hypotheses under investigation. This first part includes computer programs for computation of estimates and conducting statistical tests. It also throws light on some aspects of special experiments.

Moreover, this first part is composed of three chapters, the first of which is concerned with a general mathematical approach to the analysis of block experiments, the second deals with properties and methods of designing block experiments, and the last chapter deals with spoilt experiments.

In the second part emphasis is put on two-way, threeway and multiple-way classified experiments. We introduce basis of designs and methods of analysis.

New contribution in hyper cubic latin designs is also introduced. The second part is composed of three chapters, the first of which considers two-way classified experiments, the second three-way, and the third multiple classifactory designs. The author introduces some basis of designs of his experiments and their methods of analysis and tests of hypotheses involved. Computer programs are also supplied in this part.

In the third part emphasis is being put on carry over designs using latin squares. Methods of estimations of carry over effects are introduced as well as the statistical tests of hypotheses involving carry over effects.

PART ONE

THE DESIGN AND ANALYSIS OF BLOCK EXPERIMENTS

1.1 Introduction

A block experiment is an experiment whose purpose is to estimate the effects of certain tradiments applied to some material which is grouped into relatively homogeneous blocks. In the analysis of the results of such experiments care must be taken to ensure that decrease in variability of response arising from the grouping is reflected in the treatment estimates. Experiments of this kind constitute the najority of experiments which are subject to statistical analysis.

The purpose of this part is three-fold. First, it is intended to derive the coumon forms of analysis of such experiments using matrix methods. The systematic use of this condensed notation enables the common features of as well as the differences between the various analysis used to be better appreciated. It also brings out the essential similarity of the surlysis for different experimental errangements.

The second purpose is to prepare the comptations necessary in these analysis in a form suitable for use on automatic calculating machines. Such machines, once given

their instructions, will be able to analyse any experimental data considerably more quickly than bumen computers, and with the added advantage of relieving the tedium of routine calculation. However, before this is a practical possibility, it is necessary to prepare one set of instructions which will enable the machine to analyse all possible experiments. Unless this is done the labour of devising afresh a set of instructions for each individual design would destroy the educations of the use of such machines over manual calculation.

The matrix presentation of the accepted methods of enalysis leads immediately to the pattern of calculation which can form the basis of such a set of instructions.

Obviously, the use of the calculating machines may serve the purpose but is far inferior to the use of an electronic computer.

i.e., if the size of computational work is great, we have proposed a Fortran progrems to reduce it and make it practically applicable to be encountered in practical life.

The last function of this part is to present the problem of design. It is widely recognized that the purpose of any experiment of this kind is to estimate the effects of the treatments as accurately as possible.

It will also be shown how the problems of the analysis of spoilt experiments can be tackled using the same basic pattern of calculation.

Matrix Notations Adopted

This section is devoted to a brief summary of the properties of the principal mathematical tools we shall use in the following chapters. Readers who are familiar with matrices are advised not to omit this section altogether, as although the results are all reached by elementary means, certain non-standard notations are introduced for convenience.

A vector whose elements are all units is denoted by \underline{l} . If the length (i.e. number of rows) of the vector requires indication, this is placed as a subscript in square bracket, e.g. \underline{l}_n , a square matrix with all off-diagonal elements zero is called diagonal. If the diagonal elements are arranged as a vector \underline{x} , then the diagonal matrix is denoted by \underline{X}^{δ} and the inverse of this diagonal matrix is denoted by $\underline{X}^{-\delta}$.

The main problem arising in the following chapters is the inversion of certain square symmetric matrix, we have

$$(1^{\delta} + ab)^{-1} = 1^{\delta} - a(1^{\delta} + ba)^{-1} b$$

this can be proved by direct multiplication.

The extreme case when a is a column vector (and hence b is a row vector) is very useful. In this case 1^{δ} + ba reduces to a scalar and the inversion can be computed as

an example. Suppose $a = \lambda \underline{1}$, $b = \underline{1}^s$ then if the vectors are of length n then,

$$(1^{\delta} + \lambda \underline{1} \underline{1}')^{-1} = 1^{\delta} - \frac{\lambda}{1 + \lambda^{n}} \underline{1} \underline{1}'$$

Matrices of this type constantly occur and we define the left-hand side of this equation as

$$\left\{z_{n}(\lambda)\right\}^{-1}$$
, then
$$\left\{z_{n}(\lambda)\right\}^{-1} = z_{n}\left(\frac{-\lambda}{1+\lambda n}\right)$$

Also, we must note that $|z_n(\lambda)| = 1 + n\lambda$

CHAPTER ONE

1.2- ANALYSIS OF BLOCK EXPERIMENTS

1.2.1- The general linear set-up experiments and its analysis:

Suppose that an experiment gives rise to n observations, we require to describe the distributions from which these can be regarded as a sample, we assume the following properties:-

- i) the observations are uncorrelated with each other.
- ii) they all have equally variance σ^2 .
- iii) their mean values can be expressed as known linear compounds of a set of p parameters and $p \leq n$.
 - iv) no mean can be expressed as a linear compound of less than p other means.

Suppose we have a vector \underline{y} denotes the observations and that $\underline{\theta}$ is the vector formed from the parameters.

Then the linear sat-up is given by the equations

$$E(\underline{y}) = 3 \underline{\theta}, \quad V(\underline{y}) = \sigma^2 1^{\delta}$$

where a is n x p matrix of rank p, of known coefficients in the various linear compounds denoted by the design matrix, and σ^2 is the variance of any individual observation. The fundamental result for the whole theory of design of experiments as developed here is the Gauss-Markoff theorem, this states that:-

" the unbiased linear estimate of minimum variance of any parameter is that given by the method of least squares".

This result is so well known that its proof will not be produced here, the explicit expression for the best estimate of $\underline{\theta}$ is denoted by $\underline{\underline{\theta}}$, and

$$\hat{\underline{\Theta}} = (a,a)^{-1} a, \underline{y}$$

Suppose we write

 $\underline{x}= \vartheta \underline{\Theta} + \underline{\epsilon} \ , \ \text{taking the expectation for both}$ sides then E ($\underline{\epsilon}$) = 0

$$\underline{\theta} = (a^{\dagger} \circ)^{-1} \circ \underline{t} = (a^{\dagger} a)^{-1} a^{\dagger} (a \underline{\theta} + \underline{\epsilon})$$

$$= (\epsilon^{\dagger} a)^{-1} (a^{\dagger} a) \underline{\theta} + (a^{\dagger} a)^{-1} a^{\dagger} \underline{\epsilon}$$

$$= \underline{\theta} + (a^{\dagger} a)^{-1} a^{\dagger} \underline{\epsilon}$$

then

$$\mathbb{E}(\frac{\hat{\Theta}}{\Theta}) = \underline{\hat{\Theta}} , \text{ make } \hat{\underline{\Theta}} \text{ is an umbiased estimate for } \underline{\Theta}.$$

$$\text{also, } \hat{\underline{\Theta}} - \underline{\hat{\Theta}} = (a^{\dagger}a)^{-1} \text{ a'} \underline{\hat{\Theta}} , \underline{y} - a \underline{\hat{\Theta}} = \underline{\hat{\Theta}}$$

$$\mathbb{V}(\underline{y}) = \mathbb{E}\left\{ (\underline{y} - a \underline{\hat{\Theta}})(\underline{y} - a \underline{\hat{\Theta}})^{\dagger} \right\} = \mathbb{E}\left\{ (a^{\dagger}a)^{-1}a^{\dagger} \underline{\hat{\Theta}} \underline{\hat{\Theta}} \underline{\hat{\Theta}} - \underline{\hat{\Theta}} \right\}$$

$$= (a^{\dagger}a)^{-1}a^{\dagger} \mathbb{E}(\underline{\hat{\Theta}}, \underline{\hat{\Theta}}) \cdot \underline{\hat{\Theta}} = (a^{\dagger}a)^{-1} \underline{\hat{\Theta}} \cdot \underline{\hat{\Theta}} - \underline{\hat{\Theta}} \underline{\hat{\Theta}} - \underline{\hat{\Theta}} - \underline{\hat{\Theta}} - \underline{\hat{\Theta}} - \underline{\hat{\Theta}} \underline{\hat{\Theta}} - \underline{\hat{\Theta}}$$

Thus the variance matrix of the estimates of peremeter vector $\underline{\mathbf{e}}$ which form the centre of intrest in the design of

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determined. The usual estimate of a is determined from order of the errors in the parameter estimates can be In any actual experiment of must be estimated so that the the experiment should be recorded for use in the snelysis. To egata rgiseb Vientmilerq edt garind benimieteb xirten of the work of the saelysis. Thus the value of the variance $(s^*s)^{-1}$ must be known, and its celculation is the burden chosen experiment can be finally settled the value of involve the choice of the Before the suitability of any control of the experimenter and the problem of deaten must individual observation. This latter is often beyond the (referred to as the design matrix), and the variance of an experiments is immediately calculated from the matrix as

 $(\widehat{\Theta} \circ - \underline{\chi}) \cdot (\widehat{\Theta} \circ - \underline{\chi}) = 2$

 $= \sqrt[3]{1}^{2} - 8(8)^{2} - \sqrt[3]{1} =$

Kes X M. Z =

siter a little routine algebra

the minimized sum of squares:

where $M = L^{\delta} - s(s^*s)^{-L}$ s' is symmetric and idempotent matr:

 $\mathbb{E}(\mathbf{S}) = \text{price } (\mathbb{N} \cdot \mathbb{D}(\mathbb{X})) + \left[(\mathbb{E}(\mathbb{X})), \mathbb{N} \cdot (\mathbb{E}(\mathbb{X})) \right]$

 $O = ((X)\Xi)_*M_*((X)\Xi)$

p∈ ¤ce