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# CAIRO UNIVERSITY FACULTY OF ECONOMICS AND POLITICAL SCIENCE

# SOME METHODS OF ESTIMATION IN THE CASE OF MULTIVARIATE CAUSES OF FAILURE MODELS

By
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# SOME METHODS OF ESTIMATION IN THE CASE OF MULTIVARIATE CAUSES OF FAILURE MODELS

Ву

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TO MY GOD

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#### SUMMARY

This thesis is concerned with the problem of estimating the unknown parameters of compound and mixed models if the underlying distributions are exponential and Weibull. We shall review some statistical failure models, such as exponential and Weibull distributions for single cause of failure models and mixed and compound models for multicause of failure models. Also we shall review the methods of estimation for the compound and mixed models, and present the known properties of the estimators and some numerical investigations.

For mixed Weibull with equal shape parameter, since the Bayesian method had not been discussed previously we shall present a full Bayesian analysis for the estimation problem of the unknown parameters using complete and censored data.

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# CHAPTER (I)

# INTRODUCTION

In many practical situations in the field of life testing, we have one potential cause of failure, while in others there are more than one potential cause of failure for all items on a test. In the case of multi-causes of failure, the problem of estimating the unknown parameters for the compound and mixed models if the underlying distributions .are exponential and Weibull has been considered by many authors. For compound exponential distribution, Bayesian and non-Bayesian estimates have been obtained by Boardman and Kendell (1970), Herman and Patell (1971), Boardman (1973), Bloc and Basu (1974) and Bancorft and Dunsmore (1976), while for compound Weibull distribution, the same problem has been considered by David and Moeschberger (1971), Herman and Patell (1971), Moeschberger (1974) and Ashour, Shoukry and Mohamed (1979) ... Similarly, for mixed exponential distribution, Bayesian and non-Bayesian estimates have been obtained by Mendenhall and Hader (1958), Rider (1960), Swamy and Doss (1960), Tallis and Light (1967) and Ashour and Jones (1976) while for mixed Weibull distribution, Kao (1959), Falls (1970) and Ashour and Jones (1976) obtained non-Bayesian estimates for the unknown parameters.

In the present thesis, a comperhensive review for all methods of estimation for mixed and compound models will be presented. A compartive study between all these estimates has been considered. Some random numbers for exponential and Weibull distributions have been generated and tested and can be used to check any, iterative, procedure.

For mixed Weibull model, since Bayesian method of estimation had not been discussed previously, we shall present a full Bayesian analysis to estimate the unknown parameters for mixed Weibulk with equal shape parameters in the case of complete and censored data. Using dependent and independent priors for the the unknown parameters, the joint posterior distribution and its trivariate, bivariate, and univariate marginal densities will be derived in the case of informative and non-informative priors. These distributions may be used to obtain Bayes' estimates.

Using univariate marginal densities, squared error loss, iterative procedure, and computer facilities, a numerical example will be carried out in the case of informative and non-informative prior.

Clearly, the results in the case of mixed exponential and simple Weibull and exponential can be obtained directly

as special cases from our results. Also our results can be extended to the case of multicauses of failure (more than two causes).

This thesis consists of six chapters, chapter II consists of some definitions and notation which will be used in the present thesis, chapter III consists of some statistical failure models, chapter IV deals with the estimation problem for compound model, chapter V deals with the same problem for mixed model, while chapter VI concerns with the Bayesian estimation for mixed Weibull model.

## Definition (2.1)

 $\sum_{i}$ 

A sample is truncated if it is drawn from an incomplete population.

# (2.3) Censored Sample

Suppose that due to restrictions on time available for testing, the experimenter frequently desires to conclude the life-test after a predetermined conditions. Sampling of this type is known as censored sampling. Bassett (1964) gave the following definition

# Definition (2.2)

A sample is said to be censored if while it is drawn from complete population, some of the individual values of its members are unknown, such sample members are themselves called censored. Censored sample can be classified into two types: single and doubly censored sample.

# (2.3.1) Single Censored Sample

In single censored sample we consider the case in which we sample from complete population but either the last few observations (in which case censoring is on the right) or the first few observations (in which case censoring is on the left) are unknown. A type I censored above sample is one in which the test is terminated at specified time T, before all the items, n say, have failed or sooner if all items have

censoring occurs when the number of survivors drops to predetermined levels.

Let n designates the total sample size and k the number of sample specimens which fail and which therefore results in completely determined life spans. Suppose that censoring occurs progressively in h stages at times  $T_i$  such that  $T_i > T_{i-1}$ ,  $i=1,2,\ldots h$ , and that at the ith stage of censoring,  $T_i$  sample specimens selected randomly from the survivors at time  $T_i$  are removed from further observation. It follows that

$$n = k + \sum_{i=1}^{h} r_i$$
 (2.1)

In type I censoring, the  $T_i$  are fixed and the number of survivors at these times are random variables. In type II censoring the  $T_i$  coincide with times of failure and are random variables, whereas the number of survivors at these times are fixed. For both types the  $r_i$ 's are fixed.

# (2.3.2) Doubly Censored Sample

Consider the case in which we sample from complete population but censoring is from two sides, i.e, the first few observations and the last few observations are unknown. Thus the incomplete lifetimes from two sides designated as doubly censored observations.

There are also two types of doubly censored samples;

type I and type II; as in single censored sample. In type I censoring, the times  $T_1$  and  $T_2$  where  $T_1 < T_2$  at which censoring occurs are fixed, while  $k_1$  small and  $k_2$  large sample members unobserved are random variables. In type II consoring the smallest  $k_1$  and the largest  $k_2$  censored observations are predetermined while  $x(k_1+1)$ ,  $x(n-k_2)$  are random variables and there are  $(n-k_1-k_2)$  uncensored observations. The case of single censoring is a special case of the doubly censoring and can be obtained by putting  $k_1 = 0$  or  $k_2 = 0$ .

# (2.4) Hazard Rate

Let F(x) be the distribution function of the time to failure random variable, x, and let f(x) be its probability density function. Then the hazard rate, h(x), can be defined as:

$$h(x) = \frac{f(x)}{1 - F(x)}$$
 (2.2)

where 1-F(x) is called the raliability at time x. Barlow and Proschan (1965) gave a probabilistic interpretation for the hazard rate, namely, h(x) dx which represents the probability of failure in a finite interval of time  $(x, x + \Delta x)$  given that the age of the components is x. The hazard rate by this definition would be:

$$\frac{F(x + \Delta x) - F(x)}{1 - F(x)} \tag{2.3}$$